

When Monte Carlo and Optimization met in a Markovian dance

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A dance, why ?

To improve Monte Carlo methods targetting: $d\pi = \pi d\mu$

- The "naive" MC sampler depends on design parameters in \mathbb{R}^p or in infinite dimension θ
- Theoretical studies characterize an optimal choice of these parameters θ_* by

$$\theta_* \in \Theta \text{ s.t. } \int H(\theta, x) d\pi(x) = 0$$

or

$$\theta_* \in \operatorname{argmin}_{\theta \in \Theta} \int C(\theta, x) d\pi(x) = 0.$$

- Strategies:
 - Strategy 1: a preliminary "machinery" for the approximation of θ_* ; **then** run the MC sampler with $\theta \leftarrow \theta_*$
 - Strategy 2: learn θ and sample **concomitantly**

To make optimization methods tractable

- Intractable objective function

$$\theta \text{ s.t. } h(\theta) = 0 \quad \text{when } h \text{ is not explicit } h(\theta) = \int_{\mathcal{X}} H(\theta, x) \, d\pi_{\theta}(x)$$

or

$$\operatorname{argmin}_{\theta \in \Theta} \int_{\mathcal{X}} C(\theta, x) \, d\pi_{\theta}(x)$$

- Intractable auxiliary quantities

Ex-1 Gradient-based methods

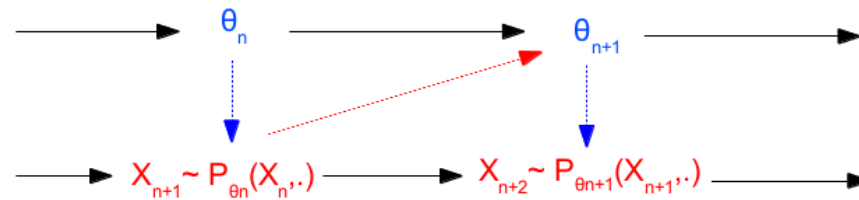
$$\nabla f(\theta) = \int_{\mathcal{X}} H(\theta, x) \, d\pi_{\theta}(x)$$

Ex-2 Majorize-Minimization methods

$$\text{at iteration } t, \quad f(\theta) \leq F_t(\theta) = \int_{\mathcal{X}} H_t(\theta, x) \, d\pi_{t, \theta}(x)$$

- Strategies: Use Monte Carlo techniques to approximate the unknown quantities

In this talk, Markov !



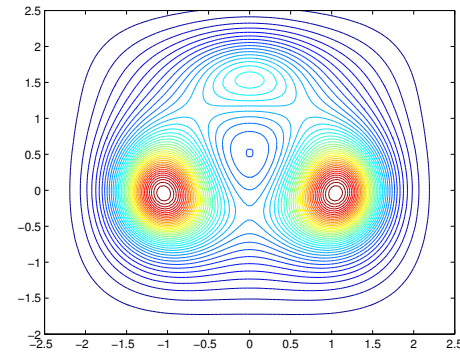
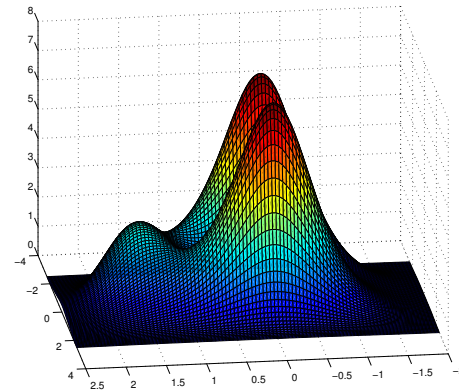
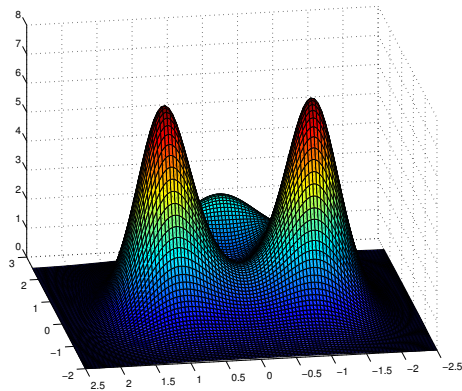
- from the Monte Carlo point of view:
which conditions on the updating scheme for convergence of the sampler ?
Case: Markov chain Monte Carlo sampler
- from the optimization point of view:
which conditions on the Monte Carlo approximation for convergence of the stochastic optimization ?
Case: Stochastic Approximation methods with Markovian inputs
- (Talk) Application to a Computational Machine Learning pbm: penalized Maximum Likelihood through Stochastic Proximal-Gradient based methods

Part I: Motivating examples

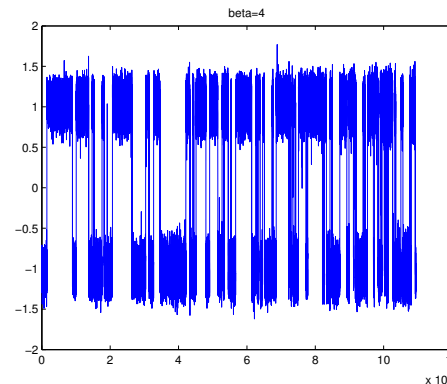
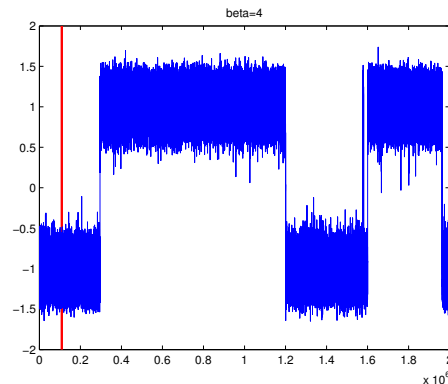
1st Ex. Adaptive Importance sampling by Wang-Landau approaches (1/6)

The problem

- A highly multimodal target density $d\pi$ on $X \subseteq \mathbb{R}^d$.



- Two samplers with different behaviors (plot: the x -path of a chain in \mathbb{R}^2)



1st Ex. (2/6)

The strategy for choosing the proposal mechanism

- A family of proposal mechanisms obtained by biasing locally the target:
 - given a partition X_1, \dots, X_I of X ,
 - for any weight vector $\theta = (\theta(1), \dots, \theta(I))$

$$d\pi_\theta(x) = \frac{1}{\sum_{i=1}^I \frac{\theta_\star(i)}{\theta(i)}} \sum_{i=1}^I 1_{X_i}(x) \frac{d\pi(x)}{\theta(i)}, \quad \text{with } \theta_\star(i) := \int_{X_i} d\pi(u).$$

- Optimal proposal: $d\pi_{\theta_\star}$ <proof>
- Unfortunately, θ_\star unavailable.

1st Ex. (3/6)

If π_{θ_\star} were available

- The algorithm would be:

- Sample X_1, \dots, X_n, \dots i.i.d. with distribution $d\pi_{\theta_\star}$ (or a MCMC with target $d\pi_{\theta_\star}$)

- Compute the importance ratio

$$\frac{d\pi}{d\pi_{\theta_\star}}(X_k) = I \sum_{i=1}^I 1_{X_i}(X_k) \theta_\star(i)$$

- When approximating an expectation, set

$$\int \phi d\pi \approx \frac{I}{T} \sum_{t=1}^T \left(\sum_{i=1}^I 1_{X_i}(X_t) \theta_\star(i) \right) \phi(X_t).$$

1st Ex. (4/6)

θ_* and therefore $d\pi_{\theta_*}$ are unknown, so ?

- $\theta_* \in \mathbb{R}^I$ collects $\int_{X_i} d\pi$ for all $i \in \{1, \dots, I\}$,

- θ_* the unique root of $\theta \mapsto \int_{X_i} H(\theta, x) d\pi_{\theta}(x) \in \mathbb{R}^I$ where for all $i \in \{1, \dots, I\}$

$$H_i(\theta, x) := \theta(i) \mathbf{1}_{X(i)}(x) - \theta(i) \sum_{j=1}^I \mathbf{1}_{X_j}(x) \theta(j).$$

thus suggesting the use of a Stochastic Approximation procedure: $\theta_* \approx \lim_t \theta_t$

$$\theta_{t+1} = \theta_t + \gamma_{t+1} H(\theta_t, X_{t+1}) \quad X_{t+1} \sim d\pi_{\theta_t}$$

- This update scheme is a normalized counter of the number of visits to X_i

<proof>

1st Ex. (5/6)

The algorithm: Wang-Landau based procedures

- Initialisation: a weight vector θ_0

Repeat for $t = 1, \dots, T$

- sample a point $X_{t+1} \sim d\pi_{\theta_t}$
- update the estimate of θ_*

$$\theta_{t+1} = \theta_t + \gamma_{t+1} H(\theta_t, X_{t+1}) \quad .$$

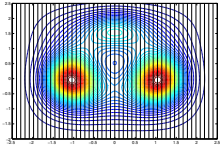
where $X_{t+1} \sim P_{\theta_t}(X_t, \cdot)$ and P_{θ} inv. wrt $d\pi_{\theta}$.

- Expected:

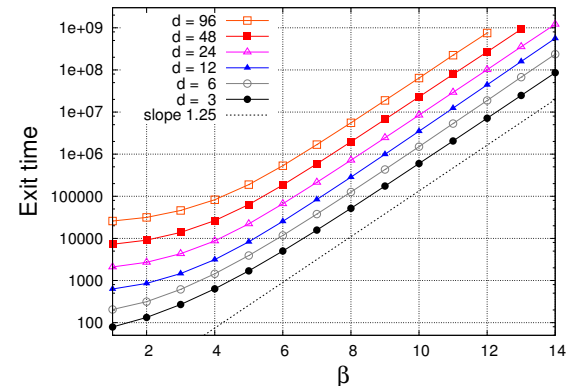
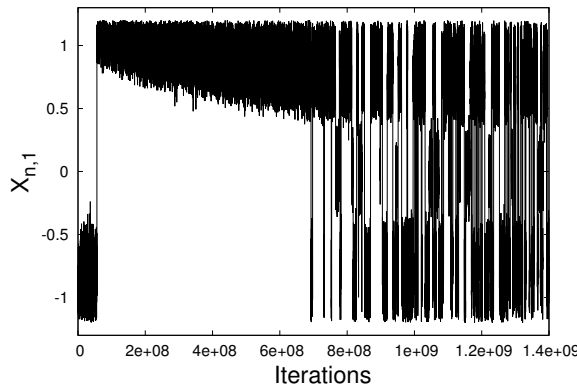
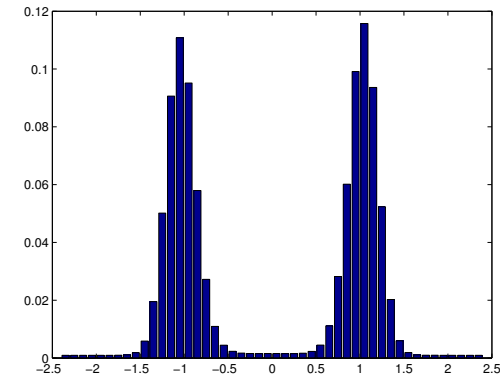
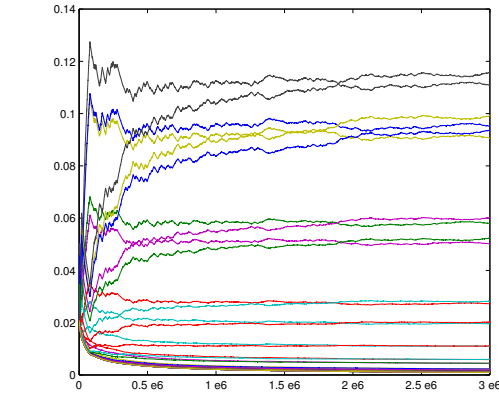
- the convergence of θ_t to θ_* : SA scheme, fed with adaptive (controlled) MCMC sampler,
- the convergence of the distribution of X_t to $d\pi_{\theta_*}$

1st Ex. (6/6)

Does it work ? Plot: convergence of θ_t and first exit times from one mode



► see F, Kuhn, Jourdain, Lelièvre, Stoltz (2014); F, Jourdain, Lelièvre, Stoltz (2015,2017,2018) for studies of these Wang-Landau bases algorithms; including self-tuned SA update rules (γ_t is random).



Conclusion of the 1st example

- Iterative sampler
- Each iteration combines : (i) a sampling step $X_{t+1} \sim P_{\theta_t}(X_t, \cdot)$; and (ii) an optimization step to update the knowledge of some optimal parameter.
- The points $\{X_1, \dots, X_t, \dots\}$ can be seen as the output of a controlled Markov chain

$$\mathbb{E} [f(X_{t+1}) | \mathcal{F}_t] = P_{\theta_t}(X_t, \cdot) \quad \mathcal{F}_t := \sigma(X_{0:t}, \theta_0)$$

where P_{θ} has $d\pi_{\theta}$ as its unique invariant distribution.

- The convergence of the parameter θ_t is the convergence of a SA scheme with "controlled Markovian" dynamics

$$\theta_{t+1} = \theta_t + \gamma_{t+1} H(\theta_t, X_{t+1})$$

2nd Example: penalized ML in latent variable models (1/6)

- An example from Pharmacokinetic:
 - N patients.
 - At time 0: dose D of a drug.
 - For patient $\#i$, observations Y_{i1}, \dots, Y_{iJ_i} giving the evolution of the concentration at times t_{i1}, \dots, t_{iJ_i} .

- The model:

$$Y_{ij} = \mathcal{F}(t_{ij}, X_i) + \epsilon_{ij} \quad \epsilon_{ij} \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma^2)$$

where $X_i \in \mathbb{R}^L$ is modeled as

$$X_i = Z_i \beta + d_i \in \mathbb{R}^L \quad d_i \stackrel{i.i.d.}{\sim} \mathcal{N}_L(0, \Omega) \text{ and independent of } \epsilon_{\bullet}$$

and Z_i known matrix s.t. each row of X_i has in intercept (fixed effect) and covariates.

- Statistical analysis: (i) estimation of $\theta = (\beta, \sigma^2, \Omega)$, under sparsity constraints on β ; (ii) selection of the covariates based on $\hat{\beta}$.

2nd Ex. (2/6)

Penalized Maximum Likelihood

- The likelihood of $Y := \{Y_{ij}, 1 \leq i \leq N, 1 \leq j \leq J_i\}$ is not explicit:
 - The distribution of $Y_{i,j}$ given X_i is simple; the distribution of X_i is simple.
 - The joint distribution has an explicit expression - It is an example of latent variable model:

$$\log L(Y; \theta) = \log \int p(Y, x_{1:N}; \theta) d\nu(x_{1:N})$$

- Sparsity constraints on the parameter θ : through a penalty term $g(\theta)$
- The penalized ML is of the form

$$\operatorname{argmin}_{\Theta} (-\log L(Y; \theta) + g(\theta))$$

with an intractable objective function.

2nd Ex. (3/6)

What about first-order methods for solving the optimization ?

- On the likelihood term:

- Usually regular enough so that the Gradient exists and <proof>

$$\begin{aligned}\nabla_{\theta} \log L(Y; \theta) &= \int \frac{\partial_{\theta} p(Y, x; \theta)}{p(Y, x; \theta)} \frac{p(Y, x; \theta) d\mu(x)}{\int p(Y, z; \theta) d\mu(z)} \\ &= \int \partial_{\theta} (\log p(Y, x; \theta)) \underbrace{d\pi_{\theta}(x)}\end{aligned}$$

the a posteriori distribution of x given Y
the dep upon Y is omitted

- the a posteriori distribution is known up to a normalizing constant.

- On the penalty term

- May be non smooth, but: convex and lower semi-continuous

- Hence a Proximal operator (implicit gradient) is associated - <See the talk, on tuesday afternoon>.

2nd Ex. (4/6)

What about EM-like methods for solving the optimization ?

- Expectation-Maximization introduced to solve below: modified for a minimization

$$\operatorname{argmin}_{\theta \in \Theta} \left(\log \int_{\mathcal{X}} p(x; \theta) d\mu(x) - g(\theta) \right)$$

where the first part is untractable; by iterating two steps

- Expectation step

$$Q(\theta, \theta_t) := \int \log p(x; \theta) \frac{p(x; \theta_t) d\mu(x)}{\int p(z; \theta_t) d\mu(z)} = \int \log p(x; \theta) d\pi_{\theta_t}(x)$$

- Minimization step

$$\theta_{t+1} := \operatorname{argmin}_{\theta} (-Q(\theta, \theta_t) + g(\theta)).$$

- $\theta \mapsto Q(\theta, \theta_t)$ is an integral which is untractable; $d\pi_{\theta}$ is known up to a normalizing constant.

2nd Ex. (5/6)

- Both in EM-like approaches and in gradient-based approaches,
 - faced with untractable auxiliary quantities of the form

$$\int_{\mathcal{X}} H(\theta, x) \, d\pi_{\theta_t}(x) \tag{1}$$

at iteration t of the optimization algorithm.

- untractable integral; $d\pi_{\theta}$ is often known up to a normalizing constant.

- What kind of stochastic approximation of the integral (1) at iteration t ?

- Quadrature techniques: poor behavior w.r.t. the dimension of \mathcal{X}

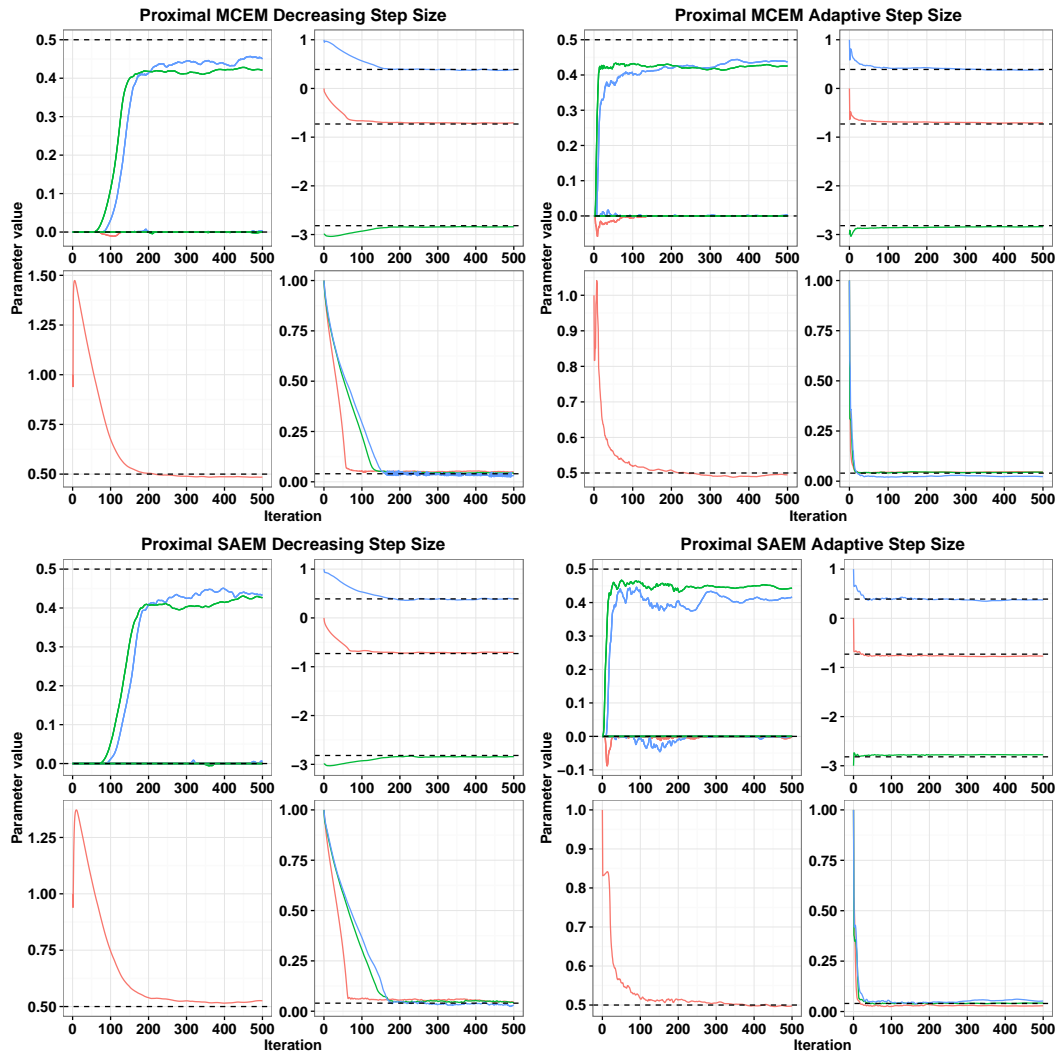
- I.i.d. samples from π_{θ_t} to define a Monte Carlo approximation: not possible, in general.

- use T samples from a MCMC sampler $\{X_{j,t+1}, j \geq 0\}$ with unique inv. dist. $d\pi_{\theta_t}$.

2nd Ex. (6/6)

Does it work ?

see F,Moulines (2003)
for EM-like approaches;
see Atchadé,F,Moulines
(2017) and
F,Ollier,Samsó (2018)
for gradient-based
approaches;
see F,Ollier,Samson
(2018) for the parallel
between EM-like
and Gradient-based
techniques



Conclusion of the 2nd example

- Iterative optimization technique

- Each iteration combines : (i) an update of the parameter; (ii) a sampling step $X_{j+1,t+1} \sim P_{\theta_t}(X_{j,t+1}, \cdot)$ to approximate auxiliary quantities.

- The convergence of $\{\theta_t\}_t$ is the convergence of a stochastically perturbed iterative optimization algorithm. At each iteration: an exact quantity $\int H(\theta, x) d\pi_{\theta_t}(x)$ is approximated by a Monte Carlo sum

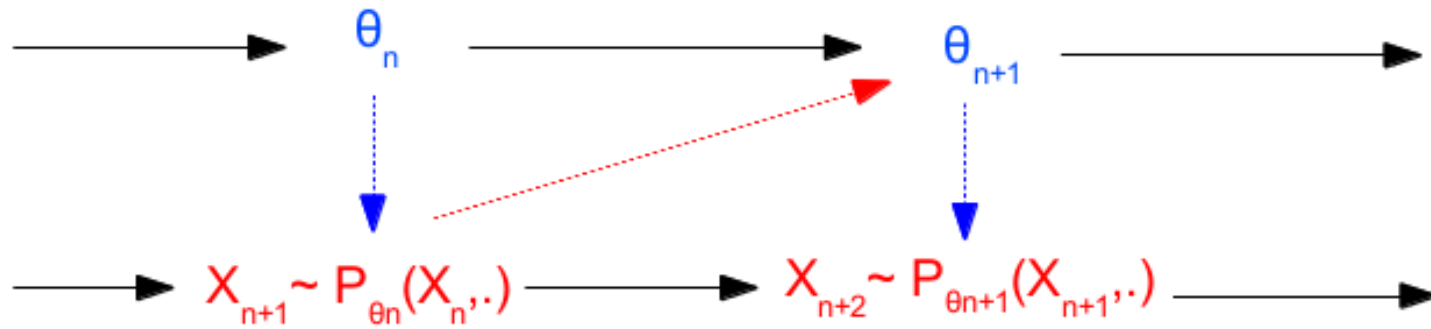
$$\int H(\theta, x) d\pi_{\theta_t}(x) \approx \frac{1}{m_{t+1}} \sum_{j=1}^{m_{t+1}} H(\theta, X_{j,t+1})$$

- The points $\{X_{j,t+1}\}_j$ satisfy

$$\mathbb{E} \left[f(X_{j,t+1}) | \mathcal{F}_t \right] = P_{\theta_t}^j(X_{0,t+1}, \cdot) \quad \mathcal{F}_t := \sigma(X_{:,0:t}, \theta_0), \quad X_{0,t+1} = X_{m_t,t}$$

where P_{θ} has $d\pi_{\theta}$ as its unique invariant distribution.

Conclusion of this first part (1/3): is a theory required ?



Conclusion of this first part (2/3): is a theory required when sampling ?

YES ! convergence can be lost by the adaption mechanism

Even in a simple case when

$$\forall \theta \in \Theta, \quad P_\theta \text{ invariant wrt } d\pi,$$

one can define a simple adaption mechanism

$$X_{t+1} | \text{past}_{1:t} \sim P_{\theta_t}(X_t, \cdot) \quad \theta_t \in \sigma(X_{1:t})$$

such that

$$\lim_t \mathbb{E} [f(X_t)] \neq \int f \, d\pi.$$

<proof> A $\{0, 1\}$ -valued chain $\{X_t\}_t$ defined by $X_{t+1} \sim P_{X_t}(X_t, \cdot)$ where the transition matrices are

$$P_0 = \begin{bmatrix} t_0 & (1-t_0) \\ (1-t_0) & t_0 \end{bmatrix} \quad P_1 = \begin{bmatrix} t_1 & (1-t_1) \\ (1-t_1) & t_1 \end{bmatrix}$$

Then P_0 and P_1 are invariant w.r.t $[1/2, 1/2]$ but $\{X_t\}$ is a Markov chain invariant w.r.t. $[t_1, t_0]$

Conclusion of this first part (3/3): is a theory required when optimizing ?

YES ! Unfortunately ,

- a biased approximation <proof>

$$\mathbb{E} \left[\frac{1}{m_{t+1}} \sum_{j=1}^{m_{t+1}} H(\theta, X_{j,t+1}) \middle| \mathcal{F}_t \right] = ? \neq \int_{\mathcal{X}} H(\theta, x) d\pi_{\theta_t}(x)$$

- For a reduced computational cost: a bias which we would like NOT vanishing i.e. $m_t = m(= 1)$.

Ex. Stochastic Approximation with controlled Markovian dynamics

$$\begin{aligned} \theta_{t+1} &= \theta_t + \gamma_{t+1} H(\theta_t, X_{t+1}) & X_{t+1} &\sim P_{\theta_t}(X_t, \cdot) \\ &= \theta_t + \gamma_{t+1} \underbrace{\int H(\theta_t, x) d\pi_{\theta_t}(x)}_{h(\theta_t)} + \gamma_{t+1} \underbrace{\left(H(X_{t+1}, \theta_t) - h(\theta_t) \right)}_{\text{non centered}} \end{aligned}$$