

Exc. 1

1) $G_X(t) = E(t^X) = \sum_{k=0}^{\infty} t^k p_k \quad (*)$

b) $\sum_{k=0}^{\infty} |t^k p_k| \leq \sum_{k=0}^{\infty} |t|^k < \infty$ pour tout $t \in]-1, 1[$.

• pour $t=1$ ou $t=-1$:

$$\sum_{k=0}^{\infty} |t^k p_k| = \sum_{k=0}^{\infty} p_k = 1 < \infty$$

↳ la série (*) est absolument convergente $\forall t \in [-1, 1]$.

↳ G_X est bien définie sur $[-1, 1]$.

c) $G'_X(t) = \sum_{k=1}^{\infty} k t^{k-1} p_k = p_1 + 2t p_2 + 3t^2 p_3 + \dots$

⇒ $G'_X(0) = p_1$

thm. de dérivation
des séries entières

$$G''_X(t) = \sum_{k=2}^{\infty} k(k-1) t^{k-2} p_k$$

⇒ $G''_X(0) = 2 p_2$

la dérivée d'ordre i

$G_X(0) = p_0$

$$G_X^{(i)}(t) = \sum_{k=i}^{\infty} k(k-1)\dots(k-i+1) t^{k-i} p_k$$

On remarque
(peut montrer) que

⇒ $G_X^{(i)}(0) = i(i-1)\dots(i-i+1) p_i = i! p_i$

⇒ $p_k = \frac{G_X^{(k)}(0)}{k!}, \quad \forall k \geq 0.$

$$d) G'_X(1) = \sum_{k=1}^{\infty} k p_k = E(X) \quad \left(= \sum_{k=0}^{\infty} k p_k \right)$$

$$G''_X(1) = \sum_{k=2}^{\infty} k(k-1) p_k = \sum_{k=1}^{\infty} k(k-1) p_k = \sum_{k=1}^{\infty} k^2 p_k - \sum_{k=1}^{\infty} k p_k$$

$$= \sum_{k=0}^{\infty} k^2 p_k - \sum_{k=0}^{\infty} k p_k = E(X^2) - E(X)$$

$$\Rightarrow \text{Var}(X) = E(X^2) - (E(X))^2 = G''_X(1) + G'_X(1) - (G'_X(1))^2$$

$$e) G_{X+Y}(t) = E(t^{X+Y}) = E(t^X \cdot t^Y) = E(t^X) E(t^Y) = G_X(t) G_Y(t)$$

\nearrow
 X, Y indep.

$$2) a) X \sim \mathcal{P}_0(\lambda) \Rightarrow p_k = \frac{\lambda^k}{k!} e^{-\lambda}, \quad \forall k \geq 0$$

$$\Rightarrow G_X(t) = \sum_{k=0}^{\infty} t^k \cdot \frac{\lambda^k}{k!} e^{-\lambda} = \sum_{k=0}^{\infty} \frac{(t\lambda)^k}{k!} e^{-\lambda}$$

$$= e^{t\lambda} \cdot e^{-\lambda} = \underline{e^{(t-1)\lambda}}, \quad \forall t \in \mathbb{R}$$

$$b) E(X) = G'_X(1)$$

$$G'_X(t) = \lambda e^{(t-1)\lambda} \Rightarrow G'_X(1) = \lambda \Rightarrow \underline{E(X) = \lambda}$$

$$G''_X(t) = \lambda^2 e^{(t-1)\lambda} \Rightarrow G''_X(1) = \lambda^2$$

$$\Rightarrow \underline{\text{Var}(X)} = G'_X(1) + G''_X(1) - (G'_X(1))^2$$

$$= \lambda + \lambda^2 - \lambda^2 = \underline{\lambda}$$

$$c) G_X^{(k)}(t) = \lambda^k e^{(t-1)\lambda} \Rightarrow G_X^{(k)}(0) = \lambda^k e^{-\lambda}$$

$$\Rightarrow \mathbb{P}(X=k) = \frac{G_X^{(k)}(0)}{k!} = \underline{\frac{\lambda^k}{k!} e^{-\lambda}}$$

$$d) G_{X+Y}(t) = G_X(t) G_Y(t) = e^{(t-1)\lambda} \cdot e^{(t-1)\mu} = e^{(t-1)(\lambda+\mu)}$$

\nearrow
 X, Y indep.

$$\Rightarrow \underline{X+Y \sim \mathcal{P}_0(\lambda+\mu)}$$

Exc. 2

$$1) \quad y \in [0, 1[\quad , \quad \mathbb{P}(T \leq y \mid T < 1) = \frac{\mathbb{P}(T \leq y \cap T < 1)}{\mathbb{P}(T < 1)}$$

$$= \frac{\mathbb{P}(T \leq y)}{\mathbb{P}(T < 1)} = \frac{F_T(y)}{F_T(1)} = \frac{1 - e^{-\lambda y}}{1 - e^{-\lambda}}$$

$$2) \quad \{X = k\} = \{[T] = k-1\} = \{k-1 \leq T < k\}$$

$$\Rightarrow \mathbb{P}(X = k) = \mathbb{P}(k-1 \leq T < k) = \int_{k-1}^k f_T(t) dt$$

$$= \int_{k-1}^k \lambda e^{-\lambda t} dt = \left[-e^{-\lambda t} \right]_{k-1}^k = e^{-\lambda(k-1)} - e^{-\lambda k}$$

$$\left(\text{ou } \mathbb{P}(X = k) = \mathbb{P}(k-1 \leq T < k) = F_T(k) - F_T(k-1) \dots \right)$$

$$\Rightarrow \mathbb{P}(X = k) = e^{-\lambda(k-1)} (1 - e^{-\lambda}) = (1-p)^{k-1} \cdot p, \quad \forall k = 1, 2, \dots$$

avec $p = 1 - e^{-\lambda}$.

$X \sim \text{Géométrique}(1 - e^{-\lambda})$

$$3) \quad \mathbb{P}(Y \leq y \cap X = k) = \mathbb{P}(\underbrace{T - [T]}_{\substack{\text{la} \\ \text{partie} \\ \text{fractionnaire} \\ \text{de } T}} \leq y \cap [T] = k-1)$$

$$= \mathbb{P}(T \leq k-1 + y \cap [T] = k-1)$$

$$= \mathbb{P}(k-1 \leq T \leq k-1 + y) \quad \text{car } y \in [0, 1[$$

$$= e^{-\lambda(k-1)} - e^{-\lambda(k-1+y)} = e^{-\lambda(k-1)} (1 - e^{-\lambda y})$$

$$4) \quad F_Y(y) = \mathbb{P}(Y \leq y) = \sum_{k=1}^{\infty} \mathbb{P}(Y \leq y \cap X = k)$$

$$= \sum_{k=1}^{\infty} e^{-\lambda(k-1)} (1 - e^{-\lambda y}) = \frac{1}{1 - e^{-\lambda}} (1 - e^{-\lambda y}) \quad \left(\begin{array}{l} \text{la même} \\ \text{proba que} \\ \mathbb{P}(T \leq y \mid T < 1) \end{array} \right)$$

série géométrique de raison $e^{-\lambda}$ $\forall y \in [0, 1[$

↳ la partie fractionnaire $T - [T]$
 a la même loi que
 la loi de T sachant $\{T < 1\}$.
 conditionnelle

$$5) f_Y(y) = F_Y'(y) = \frac{\lambda}{1-e^{-\lambda}} e^{-\lambda y}, \quad \forall y \in [0,1].$$

$$\mathbb{E}(Y) = \int_0^1 y \cdot \frac{\lambda}{1-e^{-\lambda}} e^{-\lambda y} dy = \frac{1}{1-e^{-\lambda}} \int_0^1 y (-e^{-\lambda y})' dy$$

$$\text{(IPP)} \rightsquigarrow = \frac{1}{1-e^{-\lambda}} \left([y e^{-\lambda y}]_0^1 + \int_0^1 e^{-\lambda y} dy \right)$$

$$= \frac{1}{1-e^{-\lambda}} \left(-e^{-\lambda} - \left[\frac{e^{-\lambda y}}{\lambda} \right]_0^1 \right)$$

$$= \frac{1}{1-e^{-\lambda}} \left(-e^{-\lambda} - \frac{e^{-\lambda}}{\lambda} + \frac{1}{\lambda} \right)$$

$$= \frac{1}{\lambda} - \frac{e^{-\lambda}}{1-e^{-\lambda}}$$

• Une autre méthode : $Y = T - [T] = T - X + 1$

$$\Rightarrow \mathbb{E}(Y) = \mathbb{E}(T) - \mathbb{E}(X) + 1$$

$$= \frac{1}{\lambda} - \frac{1}{1-e^{-\lambda}} + 1 = \frac{1}{\lambda} - \frac{e^{-\lambda}}{1-e^{-\lambda}}$$

car $X \sim \text{Geom}(1-e^{-\lambda})$.

$$\text{Var}(Y) = \mathbb{E}(Y^2) - (\mathbb{E}(Y))^2 = \frac{2}{\lambda^2} - \frac{e^{-\lambda}}{1-e^{-\lambda}} \left(\frac{2}{\lambda} + 1 \right) - \left(\frac{1}{\lambda} - \frac{e^{-\lambda}}{1-e^{-\lambda}} \right)^2 = \frac{1}{\lambda^2} - \frac{e^{-\lambda}}{(1-e^{-\lambda})^2}$$

$$\text{car } \mathbb{E}(Y^2) = \int_0^1 y^2 \cdot \frac{\lambda}{1-e^{-\lambda}} e^{-\lambda y} dy = \frac{1}{1-e^{-\lambda}} \int_0^1 y^2 (-e^{-\lambda y})' dy$$

$$\text{(IPP)} = \frac{1}{1-e^{-\lambda}} \left([y^2 e^{-\lambda y}]_0^1 + 2 \int_0^1 y e^{-\lambda y} dy \right) = \frac{-e^{-\lambda}}{1-e^{-\lambda}} + \frac{2}{\lambda} \mathbb{E}(Y) = \frac{2}{\lambda^2} - \frac{e^{-\lambda}}{1-e^{-\lambda}} \left(\frac{2}{\lambda} + 1 \right)$$