

Statistical Prediction of Road Traffic. From Data to Models and vice versa

Fabrice Gamboa

Collaboration with: Guillaume Allain, Philippe Goudal, Jean-Michel Loubes, Jean-Noël Kien

ENBIS-5th of September



Overview

- 1 Collaboration overview
- 2 Industrial context
- 3 Road traffic models : examples
- 4 Upstream research

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The actors of the collaboration

- **Mediamobile Vtrafic** 

Mediamobile ensures the production and broadcasting of reliable and pertinent real-time traffic information. Founded in 1996, Mediamobile originated from a partnership between TDF Group (European leader in media content broadcast) and the automotive manufacturer Renault in the framework of a European Program for Research and Development of Intelligent Transportation.

- **Institut de Mathématiques de Toulouse**



The Toulouse Mathematics Institute, CNRS Research Laboratory, federates the mathematics community of the Toulouse area. One of the biggest mathematical team in France (around 400 people)

People involved in the collaboration

Six years collaboration leading to three patents. Actual people involved

- Mediamobile Vtrafic 
 - Philippe Goudal head of the prediction department
 - Guillaume Allain Engineer has been Engineer/CIFRE Ph. D Student of the project
 - Jean-Noël Kien Engineer/CIFRE Ph. D Student
- Institut de Mathématiques de Toulouse 
 - Fabrice Gamboa Professor
 - Jean-Michel Loubes Professor
 - Elie Maza Associated Professor
 - Jean-Noël Kien Engineer/CIFRE Ph. D Student
 - Thibault Espinasse Ph. D Student

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Mediamobile's task

- **Gathering** raw traffic information
- **Processing** and agregating
- **Broadcasting** (radio, www, mobile device...)

⇒ Fancy new services : **forecasting** and **dynamic routine**

Industrial constraints :

- **coverage** { each road of the network
from real time to long run
- **quality/accuracy** { controlled speed prediction error
controlled jam prediction error
- **user friendly** { automatable
adaptative
easy to update

Road traffic data-Road network

What is a road network ?

- Graph composed of a set of pair (**edges, vectices**)
- **Complexity** of the graph → *Functional Road Classes (FRC)*
- **FRC** → road type classification (arterial, collector, local road...)

FRC	Number of edges	$\sum L[\text{km}]$
0	46 175	22 580
1	232 572	42 793
2	462 907	75 453
3	998 808	175 790
{0,1,2,3}	1 740 462	316 616

Tab: Number of edges by FRC

- Network coverage depends on the FRC

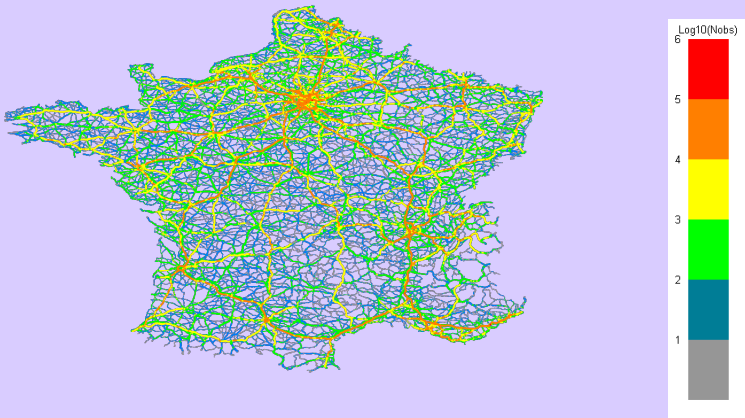


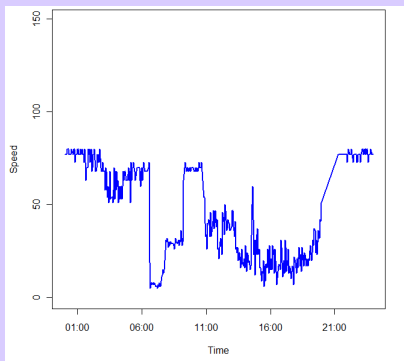
Fig: Network coverage by all FRC {0,1,2,3} from 03/01/2009 to 05/31/2009

Speed data

What is a speed data ?

Loop sensor

- speed calculated from flow and density (conservation law)



Pros

- More accurate
- 3min constant frequency

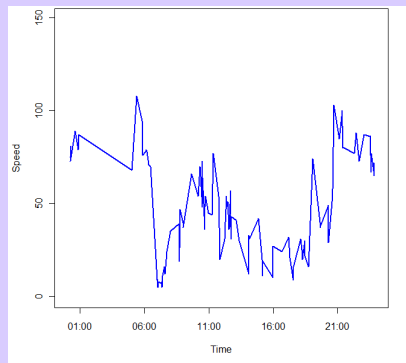
Cons

- Located only in main roads
- Thresholded at national speed limits

Speed data

GPS sensor : Floating Car Data

- positions are mapped on a graph → building speeds



Pros

- Can potentially cover all the graph
- Raw source of data

Cons

- Less accurate → GPS and map-matching error
- More variable → outlier emergence
- Random frequency → user feedback

Overview

1 Collaboration overview

2 Industrial context

3 Road traffic models : examples

- Sparse model to forecast
- Punctual model of road traffic
 - Calendar influence
 - Model by classification
 - How to use the observed speed of the day ?
 - Aggregation : statistical learning

4 Upstream research

Local road trafficking forecasting with l^1

Our Goal

- Approach the road traffic dynamic with local statistical models

$$V(s_q, t_{p+h}) = F(Q(s, t), \rho(s, t)...) \rightarrow V_{q,p+h} = g_{q,p,h}(\underbrace{\{V_{i,k}; i \in G, k \in T\}}_X)$$

Problems

- High dimension of X
- All $V_{i,k}$ not influent

Solution

- Regularization
- Selection

Modelizing traffic dynamic with significative effects **only**

$$V_{q,p+h} = g_{q,p,h}(V_{i,k}) \rightarrow V_{q,p+h} = \sum_{i \in G, k \in T} \beta_{i,k} \cdot V_{i,k}$$

where

$$\widehat{\beta}_{i,k} = \underbrace{K((i, k), (q, p + h))}_{\text{Kernel}}$$

Kernel selection : fit road traffic dynamic

- learning a sparse set of influence parameters

$$\widehat{\beta} = \arg \min_{\beta} \left(\|V_{q,p+h} - \sum_{i \in G, k \in T} \beta_{i,k} \cdot V_{i,k}\|^2 + \lambda \sum |\beta_{i,k}| \right)$$

Conclusion

- Short run local model
- Forecast and complete missing data
- Time and spatial road traffic dynamic used

Improve accuracy of short/long run predictions with weather data

Partnership between Mediamobile and Météo-France

Rupture model

$$V(x, t_1) = V(x, t_0) + C(.) \times \mathbb{1}_{M(x, t_1) \neq M(x, t_0)} \quad \text{with } t_1 - t_0 < \tau_{sta}$$

- τ_{sta} : timespan for a stationary traffic flow

$C(.)$ correction term can depend on :

- edge x : road specifications, geographical areas
- nature and intensity of the weather evolution
- traffic state at t_0 : $V(x, t_0)$

Model selection based on $C(\cdot)$ structure

Linear thresholded bias model

If $V(x, t_0) \geq \alpha$,

$$V(x, t_1) = V(x, t_0) - \underbrace{\beta}_{\text{correction term}} \cdot (V(x, t_0) - \underbrace{\alpha}_{\text{break parameter}})$$

Or else,

$$V(x, t_1) = V(x, t_0)$$

Advantages

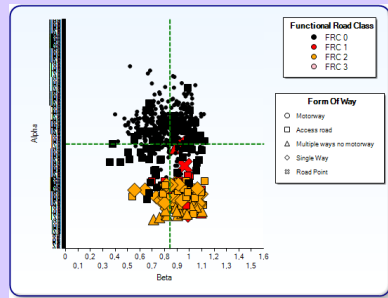
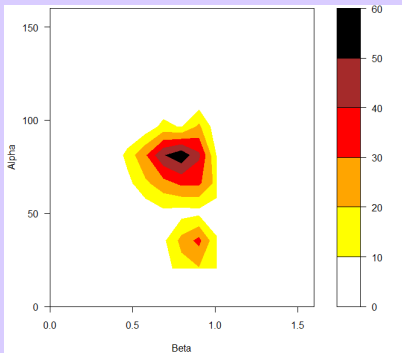
- takes traffic state into consideration
- thresholded model yields interpretable model

Drawback

- edge by edge model

Network generalization

Repartition of (α, β) parameters



Results

- β can be generalized
- Repartition of α depends on the FRC

Results with 10 000 random edges FRC 0

Weather condition	α (% of FreeFlowSpeed)	β	# obs
Low rain	93%	0,95	7236
Medium and Strong rain	90%	0.95	3316
Freezing rain	NA	NA	0
Rain and Snow	94%	0.97	1011
Snow	83%	0.96	2621
Hail	NA	NA	0
Drizzle	89%	0.90	615

For instance, let the free flow speed equals 100 km/h : a car travels at 130 km/h on a freeway and strong rain appears.

Since $130 > 90\% \cdot 100$, car speed **decreases** to $130 - 95\% \cdot (130 - 90) =$
92 km/h.

Model the relationship between speeds and calendar

How it is used :

- $D \neq$ day of the prediction
 - the speed curve is not observed
 - « Inboard configuration » \Rightarrow low **complexity**
- $$\left. \begin{array}{l} \rightarrow D \neq \text{day of the prediction} \\ \rightarrow \text{the speed curve is not observed} \end{array} \right\} \forall p, h \quad X = C$$

mathematical model : linear model with k fixed

$$g(t_k, x) = \beta_0 + \left. \begin{array}{l} \beta_1 \mathbb{1}_{\{c=\text{Monday}\}} + \beta_2 \dots \\ \beta_8 \mathbb{1}_{\{c=\text{January}\}} + \beta_{19} \dots \\ \beta_{20} \mathbb{1}_{\{c=\text{Hollidays}\}} + \dots \end{array} \right\} \text{one oder effects}$$

$$+ \left. \begin{array}{l} \beta_{1,8} \mathbb{1}_{\{c=\text{Monday} \cap \text{January}\}} + \dots \\ \beta_{1,20} \mathbb{1}_{\{c=\text{Monday} \cap \text{Hollidays}\}} + \dots \\ \dots \end{array} \right\} \text{Second order effects}$$

2

Drawbacks :

- Functionnal aspects are lost
- $(N + 1) \times K$ effects

Model by classification

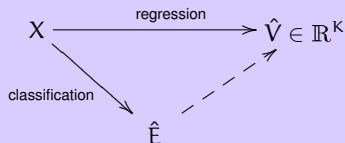
- **Functional mixture model**
- speed curve V is represented as a **finite number of patterns** :

$$f_1, \dots, f_i, \dots, f_m \text{ avec } f_i \in \mathbb{R}^K$$

$$V = \sum_{i=1}^m \mathbb{1}_{E=i} f_i + \epsilon_i \text{ et } f^* = f_E$$

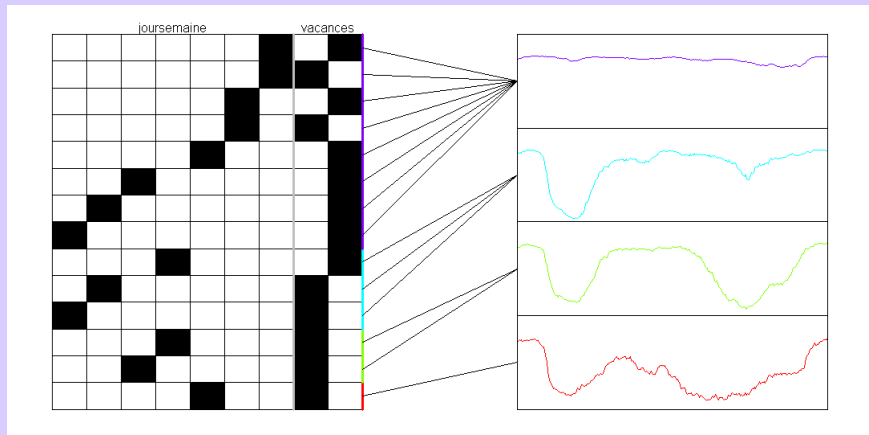
$$\left. \begin{array}{l} E \in \{1, \dots, m\} \text{ i.i.d. hidden R.V.} \\ \epsilon_i \in \mathbb{R}^K, \epsilon_i \sim \mathcal{N}(0, \Sigma_i \in \mathcal{M}_{K,K}) \end{array} \right\} \mathbb{E}[V|E=i] = f_i, \text{ Var}[V|E=i] = \Sigma_i$$

- **Classification** of E then prediction of V by f^* :



The classification model

$X = \{\text{Day of the week, Hollidays}\}$ and $m = 4$



Model the information contained in the speed of the day

Frame :

- Prediction **in the day** D
 - Speeds V^p are **known**
- } p fixed, $X = (V^p, C)$

X Time series

- How many patterns ?
 - h big et p small :
 - ⇒ m small
 - h small and p big :
 - ⇒ m **big**

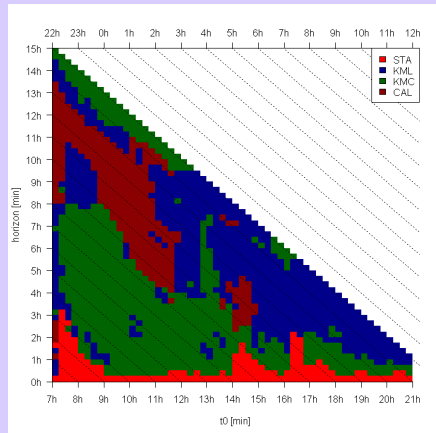
Industrial implementation

Restriction of the forecast profile :

- STA : $g(V^p) = V_p$
- KMC10 : $X = V^p$, $m = 10$ et $\tau = 1h$
- KML4 : $m = 4$ et $\tau = \infty$
- CAL4 : $X = C$, $m = 4$

Avantages :

- Hight stability
- Small processing time



Prediction of the travel time

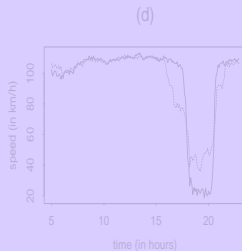
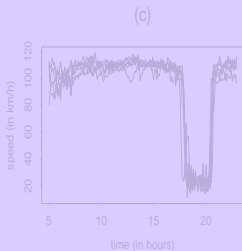
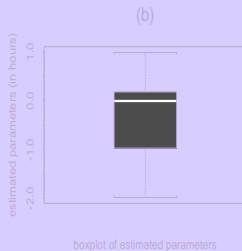
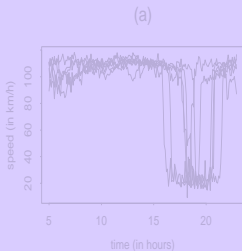
Example for a travel with $h = 1$ (forecast at one hour)

	Mean of de the relative error [%]					
	REF	STA	C10	L4	CAL4	BP
BPI (14km)	32.3	21.2	14.2	15.3	17.5	14.4
BPE (21km)	41.8	24.6	17.6	18.9	21.6	17.1
A86ES (22km)	20.4	14.7	15.4	13.2	12.6	10.1
N118W (26km)	25.4	16.7	9.6	9.8	14.3	9
A4W (35km)	21.3	17.5	12.8	13	15	11.8

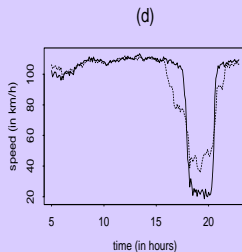
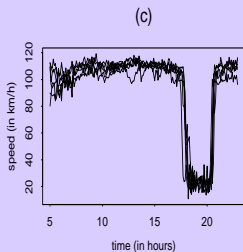
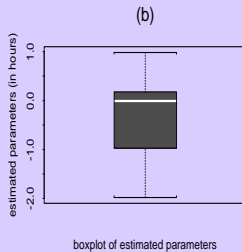
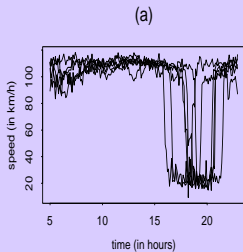
Overview

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 - Optimal estimation of a translation
 - Shift on traffic jams
 - The translated regression model
 - A more general model : Shape invariant model (SIM)
 - Model and estimation method
 - An artificial data example
 - Whittle approximation for a Gaussian field on a graph
 - General frame
 - Maximum likelihood

Shift on traffic jams



Shift on traffic jams



The translated regression model

$$Y_{i,j} = f^*(x_i - \theta_j^*) + \varepsilon_{ij}, i = 1, \dots, N, j = 1, \dots, J. \quad (1)$$

- f^* is an unknown T -periodic function
- $(\theta_j^*)_{j=1\dots J}$ is an unknown parameter of \mathbb{R}^J
- (ε_{ij}) is a Gaussian white noise with variance σ^2

Aim : Statistical inference on $(\theta_j^*)_{j=1\dots J}$, when f^* is unknown

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A more general model : Shape invariant model (SIM)

$$Y_{ij} = f_j^*(x_i) + \varepsilon_{ij} \quad i = 1 \dots n_j, j = 1 \dots J.$$

- ε is as before a Gaussian white noise with variance σ^2
- $\exists f^* : \mathbb{R} \rightarrow \mathbb{R}$ with

$$f_j^*(\cdot) = a_j^* f^*(\cdot - \theta_j^*) + v_j^* \quad (\theta_j^*, a_j^*, v_j^*) \in \mathbb{R}^3, \forall j = 1 \dots J.$$

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Model

Recall the model

$$Y_{i,j} = f^*(x_i - \theta_j^*) + \varepsilon_{ij}, i = 1, \dots, N, j = 1, \dots, J. \quad (2)$$

- f^* is an unknown T -periodic function
- $(\theta_j^*)_{j=1\dots J}$ is an unknown parameter of \mathbb{R}^J
- The design is uniform : $x_i = 2i\pi/T, i = 1, \dots, N$
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The model is not well posed. Identifiability problem

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Identifiability

Set $\alpha_j^* = \frac{2\pi}{T}\theta_j^*$.

Replacing

- α^* by $\alpha^* + c\mathbf{1} + 2k\pi$ ($c \in \mathbb{R}, k \in \mathbb{Z}^J$) (3)
- f^* by $f^*(\cdot - c)$

the observation equation remains invariant

Identifiability constraints

- Parameter set A is compact
- $\alpha^* \in A$
- If $\alpha \in A$ and $\alpha \stackrel{(3)}{=} \alpha^*$ holds then $\alpha = \alpha^*$

Examples

$$A_1 = \{\alpha \in [-\pi, \pi]^J : \alpha_1 = 0\}$$

$$A_2 = \{\alpha \in [-\pi, \pi]^J : \sum \alpha_j = 0 \text{ and } \alpha_1 \in [0, 2\pi/J]\}$$

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Estimation procedure

Main simple idea

For any $c \in \mathbb{R}$ the shift operator T_c defined on T -periodic functions

$$T_c(f) = f(\cdot - c)$$

has common **eigenvectors**

$$T_c[\exp(2i\pi/T \cdot)] = \exp(-2i\pi c/T) \exp(2i\pi/T \cdot)$$

More generally on a general group (here the torus), Fourier transform diagonalizes any translation operators acting on functions on the group (forward to extensions)

Rewrite the regression model using the eigenvectors

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Rewrite the regression model using the eigenvectors

Rewriting the model in terms of the Fourier transform

Taking the DFT and neglecting the (deterministic) error between the DFT and the Fourier transform. The model may be rewritten as (N is odd)

$$d_{jl} = e^{-il\alpha_j^*} c_l(f^*) + w_{jl}, \quad l = -(N-1)/2, \dots, (N-1)/2, \quad j = 1, \dots, J$$

- $c_l(f^*)$ is the Fourier coefficient of f^*
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Building a M-function

- Re phased Fourier coefficients

$$\tilde{c}_{jl}(\alpha) = e^{il\alpha_j} d_{jl} \quad (\alpha \in \Lambda)$$

- Mean of Re phased Fourier coefficients

$$\hat{c}_l(\alpha) = \frac{1}{J} \sum_{j=1}^J \tilde{c}_{jl}(\alpha)$$

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Idea : The deviation $\tilde{c}_{j_l}(\alpha) - \hat{c}_l(\alpha)$ should be small for $\alpha = \alpha^*$

$$M_n(\alpha) := \frac{1}{J} \sum_{j=1}^J \sum_{-(N-1)/2}^{(N-1)/2} \delta_l^2 |\tilde{c}_{j_l}(\alpha) - \hat{c}_l(\alpha)|^2$$

- (δ_l) is l^2 sequence of weights discussed later

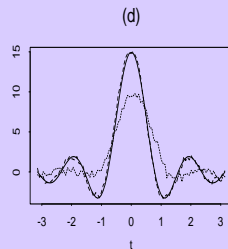
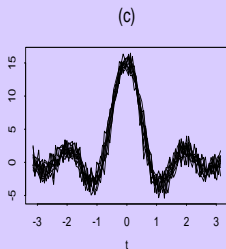
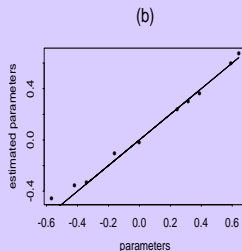
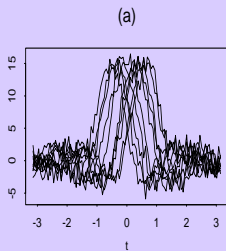
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An artificial data example



Gaussian Process on Graph : Origin of the Problem

Traffic : Predict the speed of the vehicles with missing values

For now : Spatial dependency is not exploited

Aims

- Give a model that uses spatial dependency
- Estimate the spatial correlation
- Spatial filtering

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Graph

Model : Speed process $(X_i)_{i \in G}$ indexed by the vertices G of a graph G .

Definition (Unoriented weighed graph)

$G = (G, W)$:

- G set of vertices (infinite countable)
- $W \in [-1, 1]^{G \times G}$ Weighed adjacency operator (symmetric)

Neighbors : $i \sim j$ if $W_{ij} \neq 0$

Degree of a vertex : $D_i = \#\{j, i \sim j\}$.

H_0

- $D := \sup_{i \in G} D_i < +\infty$, G has bounded degree
- $\forall i \in G, \sum_{j \in G} |W_{ij}| \leq 1$ even renormalize

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Remark :

- Our work robust to renormalization
- For \mathbb{Z} , for instance : $W_{ij}^{(\mathbb{Z})} = \frac{1}{2} \mathbf{1}_{|i-j|=1}$

W acts on $l^2(G)$:

$$\forall u \in l^2(G), \forall i \in G, (Wu)_i := \sum_{j \in G} W_{ij} u_j$$

Under H_0

W bounded as operator of $B_G := l^2(G) \rightarrow l^2(G)$:

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H'_0 : The entries of W belongs to a finite set

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General approach

Observation : Correlations are independent of the position and the orientation

Aim : Propose a *stationary* and *isotropic* model for covariances

$(X_i)_{i \in G}$ Gaussian, zero-mean, with covariance $K \in \mathbb{R}^{G \times G}$:

⇒ Characterized by K

Aim : Extension of time series

⇒ Construction MA with adjacency operator

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For $\mathbb{Z} : (\epsilon_n)_{n \in \mathbb{Z}}$ white noise

$$X_n = \sum_{k \in \mathbb{N}} a_k \epsilon_{n-k}$$

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Aim : Maximum Likelihood Estimation

⇒ Generalize Whittle's approximation

A few bibliography

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- \mathbb{Z}^d : X. Guyon
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X Gaussian centered process with covariance K is stationary if

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Spectral density

$$\text{If } r \in l^1, \exists f, K_{ij} = \frac{1}{2\pi} \int_{[0,2\pi]} f(t) \cos((j-i)t) dt := (T(f))_{ij}$$

Let $g, f(t) = g(\cos(t))$, As

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Identity resolution

Spectral decomposition

$$\exists E, \mathcal{M}, W = \int_{\mathcal{M}} \lambda dE(\lambda)$$

Definition (Identity resolution)

\mathcal{M} *Sigma-algebra* $E : \mathcal{M} \rightarrow \mathbb{B}_G$ such that $\forall \omega, \omega' \in \mathcal{M}$,

- ① $E(\omega)$ *self-adjoint operator*.
- ② $E(\emptyset) = 0, E(\Omega) = I$
- ③ $E(\omega \cap \omega') = E(\omega)E(\omega')$
- ④ *Si* $\omega \cap \omega' = \emptyset$, *then* $E(\omega \cup \omega') = E(\omega) + E(\omega')$

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- 4 *Si* $\omega \cap \omega' = \emptyset$, *then* $E(\omega \cup \omega') = E(\omega) + E(\omega')$

$$\forall i, j \in G, \forall \omega \in \mathcal{M}, \mu_{ij}(\omega) = E_{ij}(\omega)$$

Extension to a graph

Definition

$(X_i)_{i \in G}$ *Gaussian field with covariance K.*

$$\text{If } K = \int_{\text{Sp}(W)} g(\lambda) dE(\lambda),$$

- g *polynomial* : $MA_q^{(W)}$
- $\frac{1}{g}$ *polynomial* : $AR_p^{(W)}$...

Remarks :

- Conditions about g
- Equivalence with \mathbb{Z}
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- $\Theta \subset \mathbb{R}$ compact
- $(f_\theta)_{\theta \in \Theta}$ parametric family of densities associated to $K(f_\theta) = f_\theta(W)$
- Asymptotic on $(G_n)_{n \in \mathbb{N}}$ sequence of finite nested subgraphs
Example $G = \mathbb{Z} : G_n = [1, n]$.
- $\theta_0 \in \overset{\circ}{\Theta}$, $\mathbf{X} \sim \mathcal{N}(0, K(f_{\theta_0}))$
- We observe the restriction X_n of \mathbf{X} to G_n , cov : $K_n(f_\theta)$
- $m_n = \#G_n$

Aim : Estimate θ_0 by maximum likelihood :

$$L_n(\theta) := -\frac{1}{2} \left(m_n \log(2\pi) + \log \det (K_n(f_\theta)) + \mathbf{X}_n^T (K_n(f_\theta))^{-1} \mathbf{X}_n \right)$$

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Muito obrigado

