

Modeling and estimation for Gaussian fields indexed by graphs, application to road traffic prediction

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IMT

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Framework

Road traffic (Mediamobile) :

- Activity: Real-time prediction of traveling time
- Aim: Understand the speed process on the road traffic network
- Observations :
 - Fixed sensors: corrupted values
 - Cars fleet: unobserved areas
 - The graph is known
- Problem: Use the spatial dependency for:
 - Spatial completion
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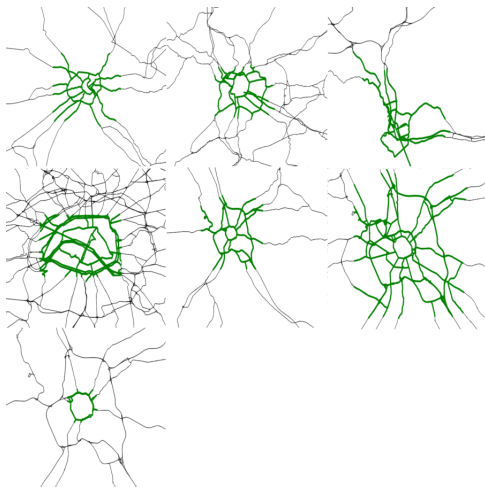
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Objective: Yield a parametric model $(\mathcal{K}_\theta)_{\theta \in \Theta}$ for covariance operators of X

Outline

- 1 Modeling
- 2 Estimation
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Speed of vehicles on the road network at a fixed time:
zero-mean Gaussian field $(X_i)_{i \in G}$ indexed by the vertices of a graph.

Aim: Chose a model for covariance operators

Modeling constraints

- Adaptability to physical modeling
- Compatibility with classical cases (time series, \mathbb{Z}^d , homogeneous tree...)
- Extension of classical tools from time series (spectral representation, Whittle's estimation...)

⇒ Define covariance operators from a spectral construction

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Graph

Model: Zero-mean Gaussian field $(X_i)_{i \in G}$ indexed by the vertices G of a graph \mathbf{G} .

Definition (Unoriented weighed graph)

$\mathbf{G} = (G, W) :$

- G set of vertices (countable)
- $W \in [-1, 1]^{G \times G}$ Weighted adjacency operator (symmetric)

Neighbors: $i \sim j$ if $W_{ij} \neq 0$

Degree of a vertex: $D_i = \#\{j, i \sim j\}$.

Assumption (H_0)

- $D := \sup_{i \in G} D_i < +\infty$, \mathbf{G} has bouded degree
- $\forall i \in G, \sum_{j \in G} |W_{ij}| \leq 1$ even renormalizing

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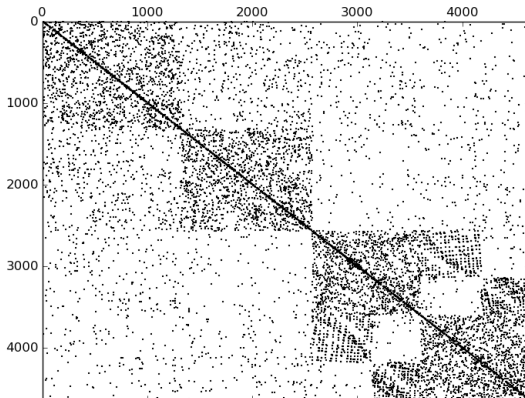
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Modeling for covariance operators

Models for covariance operators (of the speed field)

$$\mathcal{K}(f) = f(W)$$

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$$\mathcal{K}(f) = f(W)$$

W acts on $l^2(G)$:

$$\forall u \in l^2(G), \forall i \in G, (Wu)_i := \sum_{j \in G} W_{ij} u_j.$$

Modeling for covariance operators

Models for covariance operators (of the speed field)

$$\mathcal{K}(f) = f(W)$$

Under H_0

W is a bounded Hilbertian self-adjoint operator in $B_G := \mathcal{L}^2(G) \rightarrow \mathcal{L}^2(G)$:

$$\|W\|_{2,op} \leq 1.$$

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$$\mathcal{K}(f) = f(W)$$

Definition (Identity resolution)

\mathcal{M} σ -algebra $E : \mathcal{M} \rightarrow B_G$ such that $\forall \omega, \omega' \in \mathcal{M}$,

- 1 $E(\omega)$ self-adjoints projectors.
- 2 $E(\emptyset) = 0, E(\Omega) = I$
- 3 $E(\omega \cap \omega') = E(\omega)E(\omega')$
- 4 Si $\omega \cap \omega' = \emptyset$, alors $E(\omega \cup \omega') = E(\omega) + E(\omega')$

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Spectral decomposition

$$\exists E, \mathcal{M}, W = \int_{\mathcal{M}} \lambda dE(\lambda)$$

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Definition

Local measures

$$\forall i, j \in \mathbf{G}, \forall \omega \in \mathcal{M}, \mu_{ij}(\omega) = E_{ij}(\omega).$$

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$$\mathcal{K}(f) = f(W)$$

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Characterized by:

$$\forall i, j \in G, \forall k \in \mathbb{Z}, (W^k)_{ij} = \int_{\text{Sp}(W)} \lambda^k d\mu_{ij}(\lambda).$$

Models for covariance operators, spectral density

Definition (Construction of the covariance operators)

Let g be an positive function, analytic over $\text{Sp}(W)$,

$$\mathcal{K}(g) = \int_{\text{Sp}(W)} g(\lambda) dE(\lambda),$$

- g polynomial: $MA_q^{(W)}$
- $\frac{1}{g}$ polynomial: $AR_p^{(W)}$...

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Remarks:

- $\mathcal{K}(g) = g(W)$
- Dependency in W
- Analogy with \mathbb{Z}

$G = \mathbb{Z}$: compatibility with time series

Adjacency operator

$$W_{ij} = \frac{1}{2} \mathbf{1}_{|i-j|=1}.$$

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$$\forall i, j \in G, \forall k \in \mathbb{Z}, \left(W^k \right)_{ij} = \frac{1}{\pi} \int_{[-1,1]} \lambda^k \frac{T_{|j-i|}(\lambda)}{\sqrt{1-\lambda^2}} d\lambda.$$

T_k : $k^{\text{ième}}$ Chebychev polynomials

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Model

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Spectral density

$$f(t) = g(\cos(t))$$
$$\mathcal{K}(g)_{ij} = \frac{1}{2\pi} \int_{[-\pi, \pi]} f(t) \cos((j-i)t) dt := (\mathcal{T}(f))_{ij}$$

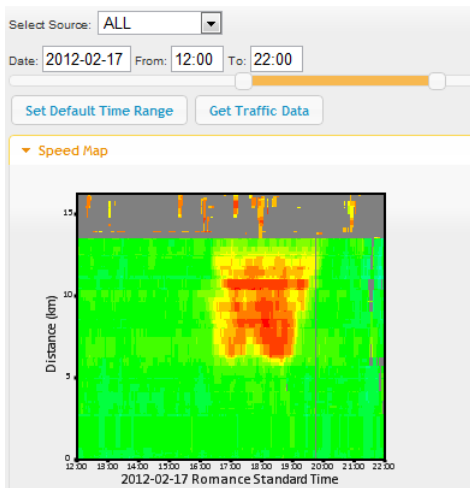
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The concrete problem



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Ideas

Framework: Parametric model of covariances operators

$$\mathcal{K}(f_\theta) = f_\theta(W).$$

Aim: Parametric estimation

Remark: Spectral density \sim Asymptotic eigendistribution of the covariance operators

Computational issues

- log det Term of the log-likelihood
- Γ^{-1} term of the log-likelihood

Other important ideas

- Trace measure
- Tapered periodogram

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Problem:

- $\Theta \subset \mathbb{R}$ compact
- $(f_\theta)_{\theta \in \Theta}$ parametric family of spectral densities associated to $\mathcal{K}(f_\theta) = f_\theta(W)$
- Asymptotic on $(\mathbf{G}_n)_{n \in \mathbb{N}}$ sequence of nested subgraphs induced by \mathbf{G}
Example $G = \mathbb{Z}$: $G_n = [1, n]$.
- $\theta_0 \in \overset{\circ}{\Theta}$, $\mathbf{X} \sim \mathcal{N}(0, \mathcal{K}(f_{\theta_0}))$
- We observe the restriction X_n of \mathbf{X} to \mathbf{G}_n , cov : $\mathcal{K}_n(f_\theta)$
- $m_n = \#\mathbf{G}_n$

Aim: Estimate θ_0 with a maximum likelihood method:

$$L_n(\theta) := -\frac{1}{2} \left(m_n \log(2\pi) + \log \det(\mathcal{K}_n(f_\theta)) + X_n^T (\mathcal{K}_n(f_\theta))^{-1} X_n \right)$$

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Computational issues: Maximize an approximation of the log-likelihood

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$$\frac{1}{n} \log \det(\mathcal{T}_n(f_\theta)) \rightarrow \frac{1}{2\pi} \int_{[0, 2\pi]} \log(f_\theta) dt.$$

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Classical case \mathbb{Z}

Computational issues: Maximize an approximation of the log-likelihood

$$\bar{L}_n(\theta) := -\frac{n}{2} \left(\log(2\pi) + \frac{1}{2\pi} \int_{[0,2\pi]} \log(f_\theta) dt + \frac{1}{n} \mathbf{X}_n^T (\mathcal{T}_n(f_\theta))^{-1} \mathbf{X}_n \right)$$

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Aim: Extension to the graph case

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Let $\delta_n = \#\delta G_n$.

Example $G = \mathbb{Z}$: $\delta_n = 2$

log det approximation for a graph

Assumption (Existence of the trace measure)

$$H_1 : \exists \mu, \frac{1}{m_n} \sum_{g \in G_n} \mu_{gg} \rightarrow \mu$$

Assumption (Edge effects)

$$H_2 : \delta_n = o(m_n)$$

Whittle's approximation for \mathbf{G} , log det

$$\frac{1}{m_n} \log \det (\mathcal{K}_n(f_\theta)) \rightarrow \int_{\text{Sp}(W)} \log (f_\theta) d\mu(t)$$

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$$\forall \mathbf{A} \in M_{m_n}(\mathbb{R}), \mathbf{b}_n(\mathbf{A}) = \frac{1}{\delta_n} \sum_{i,j \in G_n} |A_{ij}|$$

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Lemma (Asymptotic homomorphism)

$$\mathbf{b}_n\left(\mathcal{K}_n(f)\mathcal{K}_n(g) - \mathcal{K}_n(fg)\right) \leq \frac{1}{2}\alpha(f)\alpha(g),$$

where, if $f = \sum_k f_k x^k$,

$$\alpha(f) = \sum_k |f_k| (k + 1)$$

Consistency

Let $\theta_n, \bar{\theta}_n, \tilde{\theta}_n$ resp. arg max of

$$L_n(\theta) := -\frac{1}{2} \left(m_n \log(2\pi) + \log \det(\mathcal{K}_n(f_\theta)) + \mathbf{X}_n^T (\mathcal{K}_n(f_\theta))^{-1} \mathbf{X}_n \right)$$

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Assumption (H_3)

- $\theta \rightarrow f_\theta$ *injective*
- $\forall \lambda \in Sp(W), \theta \rightarrow f_\theta(\lambda)$ *continuous*.
- $\forall \theta \in \Theta, \alpha(\log(f_\theta)) \leq \rho < +\infty$

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Theorem (Consistency of the Whittle's estimate)

The estimators $\theta_n, \bar{\theta}_n, \tilde{\theta}_n$ converge $P_{f_{\theta_0}}$ -a.s. to the true value θ_0 .

Asymptotic normality and efficiency

We need:

$$\frac{1}{\sqrt{m_n}} \mathbb{E} [L'_n(\theta_0)] \rightarrow 0$$

Problem: Not true in general !!!

- \mathbb{Z}^d : X. Guyon, R. Dahlhaus
- Order: $\frac{\delta_n}{m_n}$

Solution: Extension of the tapered periodogram \mathcal{Q} .

Framework:

- **Strong** assumptions on the symmetries of the graph
- Construction of \mathcal{Q}
- AR_L
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Asymptotic normality and efficiency

Tapered likelihood

$$-2L_n^{(u)}(\theta) := m_n \log(2\pi) + m_n \int \log(f_\theta(x)) d\mu(x) + X_n^T \left(Q_n \left(\frac{1}{f_\theta} \right) \right) X_n.$$

Asymptotic normality and efficiency

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$$\theta_n^{(u)} = \arg \max L_n^{(u)}.$$

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Theorem (Asymptotic normality)

For $\theta_0 \in \mathring{\Theta}$, in the AR_L or MA_L cases, and under assumptions on the graph and the family of spectral densities, $\theta_n^{(u)}$ converges to θ_0 , and is asymptotically normal and efficient:

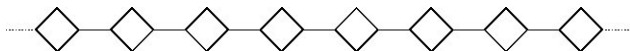
$$\sqrt{m_n}(\theta_n^{(u)} - \theta_0) \rightarrow \mathcal{N} \left(0, \left(\frac{1}{2} \int \frac{(f'_{\theta_0})^2}{f_{\theta_0}^2} d\mu \right)^{-1} \right).$$

Outline

- 1 Modeling
- 2 Estimation
- 3 Applications**

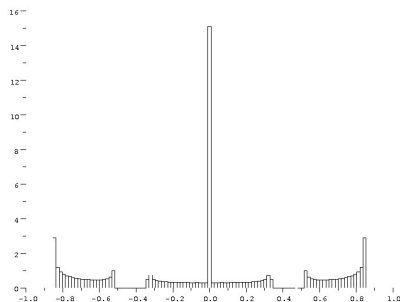
Applications

Figure: Graphe G



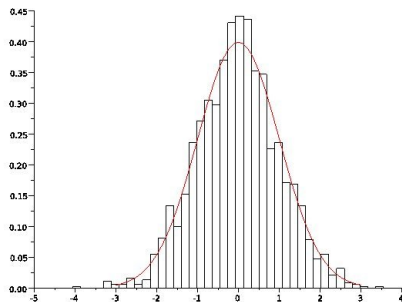
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Figure: Empirical spectral measure

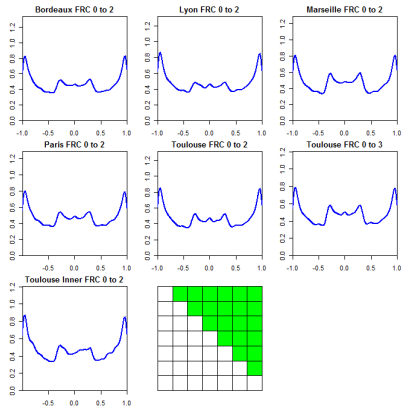
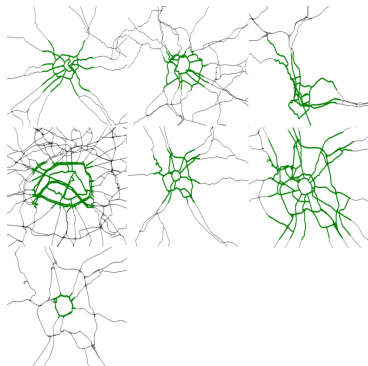


Applications

Figure: Empirical distribution of estimation error



Spectrum of the road network



Real datas

Aim: Predict missing values on *FRC* 0 in Toulouse

Real datas

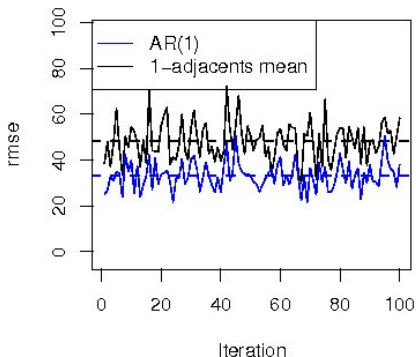
Aim: Predict missing values on *FRC* 0 in Toulouse

Protocol:

- 10% of datas hidden to test the quality of the prediction
- Model: AR_1

Real datas

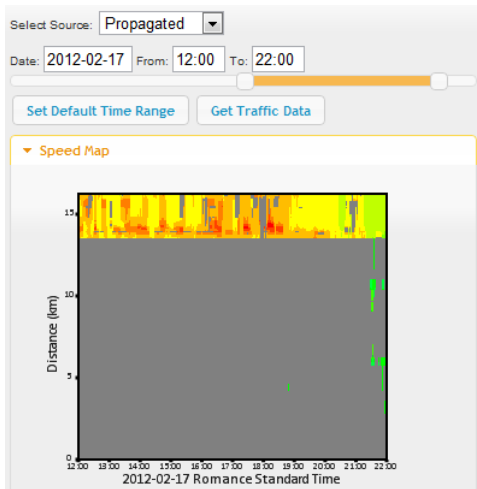
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A concrete problem



A solution ?



Projects

In progress

- Choice/estimation of the generator
- Spectral study and modeling of the road network
- Maximum likelihood for stationary processes indexed by trees
- “Blind” prediction

Future works ?

- Link with physical models
- Use approximation of manifolds by graphs
- Extension of the notion of causality

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Merci !